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Association Between Radiation Therapy, Surgery, or Observation for Localized Prostate Cancer and Patient-Reported Outcomes After 3 Years

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IMPORTANCE Understanding the adverse effects of contemporary approaches to localized prostate cancer treatment could inform shared decision making.

OBJECTIVE To compare functional outcomes and adverse effects associated with radical prostatectomy, external beam radiation therapy (EBRT), and active surveillance.

DESIGN, SETTING, AND PARTICIPANTS Prospective, population-based, cohort study involving 2550 men (\leq 80 years) diagnosed in 2011-2012 with clinical stage cT1-2, localized prostate cancer, with prostate-specific antigen levels less than 50 ng/mL, and enrolled within 6 months of diagnosis.

EXPOSURES Treatment with radical prostatectomy, EBRT, or active surveillance was ascertained within 1 year of diagnosis.

MAIN OUTCOMES AND MEASURES Patient-reported function on the 26-item Expanded Prostate Cancer Index Composite (EPIC) 36 months after enrollment. Higher domain scores (range, O-100) indicate better function. Minimum clinically important difference was defined as 10 to 12 points for sexual function, 6 for urinary incontinence, 5 for urinary irritative symptoms, 5 for bowel function, and 4 for hormonal function.

RESULTS The cohort included 2550 men (mean age, 63.8 years; 74% white, 55% had intermediate- or high-risk disease), of whom 1523 (59.7%) underwent radical prostatectomy, 598 (23.5%) EBRT, and 429 (16.8%) active surveillance. Men in the EBRT group were older (mean age, 68.1 years vs 61.5 years, P < .001) and had worse baseline sexual function (mean score, 52.3 vs 65.2, P < .001) than men in the radical prostatectomy group. At 3 years, the adjusted mean sexual domain score for radical prostatectomy decreased more than for EBRT (mean difference, -11.9 points; 95% CI, -15.1 to -8.7). The decline in sexual domain scores between EBRT and active surveillance was not clinically significant (-4.3 points; 95% CI, -9.2 to 0.7). Radical prostatectomy was associated with worse urinary incontinence than EBRT (-18.0 points; 95% CI, -20.5 to -15.4) and active surveillance (-12.7 points; 95% CI, -16.0 to -9.3) but was associated with better urinary irritative symptoms than active surveillance (5.2 points; 95% CI, 3.2 to 7.2). No clinically significant differences for bowel or hormone function were noted beyond 12 months. No differences in health-related quality of life or disease-specific survival (3 deaths) were noted (99.7%-100%).

CONCLUSIONS AND RELEVANCE In this cohort of men with localized prostate cancer, radical prostatectomy was associated with a greater decrease in sexual function and urinary incontinence than either EBRT or active surveillance after 3 years and was associated with fewer urinary irritative symptoms than active surveillance; however, no meaningful differences existed in either bowel or hormonal function beyond 12 months or in in other domains of health-related quality-of-life measures. These findings may facilitate counseling regarding the comparative harms of contemporary treatments for prostate cancer.

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Corresponding Author: Daniel A. Barocas, MD, MPH, Department of Urologic Surgery, Vanderbilt University Medical Center, A1302 Medical Center North, Nashville, TN 37232 (dan.barocas@vanderbilt.edu). he optimal management for localized prostate cancer depends on factors including risk of progression; competing risks of mortality, baseline urinary, sexual, and bowel function; and patient preferences.¹ Comparing the effectiveness and harms of radiation therapy, radical prostatectomy, and active surveillance is critical for shared decision making.² Yet comparative data have limited generalizability for several reasons, such as focusing on homogenous populations and comparing older treatments instead of contemporary robotic radical prostatectomy and intensity-modulated radiotherapy (IMRT).³⁻¹²

In this context, the Comparative Effectiveness Analysis of Surgery and Radiation (CEASAR) study, a prospective, longitudinal, population-based cohort study was developed.¹³ In light of the nearly 100% 5-year survival for men with localized prostate cancer, patient-reported disease-specific functional outcomes were selected as the primary short- and intermediate-term outcome measures. This study assessed patient-reported functional outcomes and health-related quality of life at 3 years after treatment.

Methods

The parent study accrued men diagnosed with localized prostate cancer (2011-2012) from 5 Surveillance, Epidemiology, and End Results (SEER) registries (Atlanta [Georgia], Los Angeles [California], Louisiana, New Jersey, and Utah), and the Cancer of the Prostate Strategic Urologic Research Endeavor registry. Details of the protocol have been published.¹³ Eligibility criteria were being younger than 80 years, having a prostate-specific antigen (PSA) level of less than 50 ng/mL, clinical stage T1 to T2 cancer, no nodal involvement or metastases on clinical evaluation; and being enrolled within 6 months of diagnosis (**Table**).

Patient-reported outcomes were collected via mail survey at enrollment and 6, 12, and 36 months after enrollment. If patients did not respond to 2 mailings, trained abstractor called the patient to complete the survey. A medical chart review, including clinical and treatment information, was obtained at 12 months. SEER registry data were linked to the data set. This study includes follow-up through August 2015. Institutional review board approval was obtained from each site and from Vanderbilt University Medical Center. Patients provided written informed consent.

Outcomes

The primary outcome measures were 36-month domain scores on the 26-item Expanded Prostate Cancer Index Composite (EPIC-26), a validated instrument for measuring disease-specific function.¹⁴ Domain scores range from 0 to 100, with higher score representing better function. The minimal clinically important difference (MCID), representing the magnitude of change that is clinically meaningful to patients, has been estimated for each domain using standard techniques. The distribution-based approach estimated MCID as one-third to one-half of a standard deviation, and the anchoring approach identified the magnitude

Key Points

Question What are the comparative harms of contemporary treatments for localized prostate cancer?

Findings In this prospective, population-based cohort study involving 2550 men, radical prostatectomy was associated with significant declines in sexual function compared with external beam radiation therapy (–11.9 points on a 100-point scale) and active surveillance (–16.2 points) at 3 years. Radical prostatectomy was also associated with significant declines in urinary incontinence compared with radiation and active surveillance, but there were no meaningful differences in bowel or hormonal function beyond 12 months, and no meaningful differences in health-related quality of life.

Meaning These findings may facilitate treatment counseling of men with localized prostate cancer.

of change on each domain that resulted in a change in satisfaction with treatment.¹⁵ Both techniques yielded similar MCIDs and were consistent with the a priori definition of MCID used in the power calculation of the original grant application for this study, which was one-half of a standard deviation. The sexual function domain focuses on the quality and frequency of erections (MCID, 10-12 points). The urinary incontinence (MCID, 6 points) and urinary irritative symptom (MCID, 5 points) domains ask questions about frequency; amount of urinary leakage; and dysuria, hematuria, and urinary frequency. The bowel function domain (MCID, 4 points) focuses on bowel frequency, urgency, bleeding, and pain. The hormonal domain (MCID, 4 points) assesses symptoms such as hot flashes, gynecomastia, low energy, and weight change. The baseline survey asked about pretreatment function. Previous studies have investigated the issue of recall bias for the EPIC instrument, including a study in this cohort, and adjusted differences in domain scores between those who complete the survey before treatment and those who complete it afterward range from 1.0 to 3.7 points, well below the MCID for each domain.¹⁶

Individual items from the EPIC-26 were selected a priori as secondary outcomes based on clinical relevance by content experts and patients on the study team.

Treatments were also compared with respect to healthrelated quality of life, using selected domains from the commonly used Medical Outcomes Study Short Form 36 (SF-36): physical functioning, emotional well-being, and energy and fatigue.^{17,18} Domain scores are scaled from 0 to 100, with higher scores indicating better function. The MCIDs for these domains have been estimated for patients with localized prostate cancer as 7, 6, and 9 points, respectively.¹⁹

Exposure

The main exposure was initial treatment, defined according to the following hierarchy of sources: medical chart abstraction, patient report, and SEER registry. A participant was categorized as undergoing active surveillance if this strategy was documented in the absence of treatment or if no treatment was administered within 1 year of diagnosis.

Table. Demographics and Baseline Characteristics

	No. (%) of Patlents Radical Prostatectomy (n = 1523)	External Beam Radiation Therapy (n = 598)	Active Surveillance (n = 429)	Overall (N = 2550)	P Value	
Demographics						
Age, mean (95% CI), y	61.5 (61.1-61.8)	68.1 (67.6-68.7)	66.1 (65.4-66.9)	63.8 (63.5-64.1)	<.001	
Race/ethnicity, No.	1511	597	427	2535		
White	1130 (75)	421 (71)	323 (75)	1874 (73)		
Black	187 (12)	110 (18)	61 (14)	358 (14)		
Hispanic	125 (8)	37 (6)	24 (6)	186 (7)	.02	
Asian	46 (3)	22 (4)	12 (3)	80 (3)		
Other	23 (2)	7 (1)	7 (2)	37 (1)		
Education, No.	1411	577	409	2427		
<high school<="" td=""><td>130 (9)</td><td>86 (15)</td><td>33 (8)</td><td>249 (10)</td><td></td></high>	130 (9)	86 (15)	33 (8)	249 (10)		
High school graduate	302 (21)	118 (20)	79 (19)	499 (21)		
Some college	315 (22)	133 (23)	84 (21)	532 (22)	.002	
College graduate	345 (24)	118 (20)	98 (24)	561 (23)		
Graduate or professional school	349 (24)	122 (21)	115 (28)	586 (24)		
Marital status, No.	1138	576	407	2421		
Married	1192 (83)	429 (74)	326 (80)	1947 (80)	<.001	
Comorbidity score, No.ª	1448	580	411	2439		
0-2	481 (33)	101 (17)	105 (26)	687 (28)		
3-4	624 (43)	238 (41)	162 (39)	1024 (42)	<.001	
≥5	343 (24)	241 (42)	144 (35)	728 (30)		
D'Amico prostate cancer risk, No. ^b	1521	596	427	2544		
Low risk	635 (42)	182 (31)	327 (77)	1144 (45)		
Intermediate risk	635 (42)	267 (45)	81 (19)	983 (39)		
High risk	251 (17)	147 (25)	19 (4)	417 (16)	<.001	
Prostate specific antigen, ng/mL, No.	1523	598	429	2550		
0-4	334 (22)	85 (14)	110 (26)	529 (21)		
4.1-10	1018 (67)	394 (66)	268 (62)	1680 (66)	. 001	
10.1-20	133 (9)	86 (14)	38 (9)	257 (10)	<.001	
>20	38 (2)	33 (6)	13 (3)	84 (3)		
Clinical stage, No.	1520	597	422	2539		
T1c	1140 (75)	436 (73)	357 (85)	1933 (76)	< 001	
T2	380 (25)	161 (27)	65 (15)	606 (24)	<.001	
Biopsy Gleason score, No.	1519	596	427	2542		
3 + 3	744 (49)	210 (35)	370 (87)	1324 (52)		
3 + 4	458 (30)	201 (34)	44 (10)	703 (28)	- 001	
4 + 3	170 (11)	86 (14)	7 (2)	263 (10)	<.001	
8-10	147 (10)	99 (17)	6 (1)	252 (10)		
Any hormone therapy in the first year, No.	1509	593	391	2493	·<.001	
Yes	75 (5)	265 (45)	3 (1)	343 (14)		
Survey scores						
Expanded Prostate Cancer Index Composite ^c						
Sexual domain, No. of patients	1447	558	402	2407		
Mean (95% CI)	65.2 (63.5-66.9)	52.3 (49.6-55.0)	63.1 (60.0-66.2)	61.9 (60.5-63.2)	<.001	
Urinary incontinence, No. of patients	1467	575	409	2451		
Mean (95% CI)	86.7 (85.5-87.8)	88.2 (86.7-89.6)	88.7 (87.0-90.4)	87.486.6-88.2)	.88	
Urinary irritative, No. of patients	1463	574	409	2446		
Mean (95% CI)	83.2 (82.3-84.1)	82.3 (80.9-83.7)	83.9 (82.3-85.5)	83.1 (82.4-83.8)	.10	
Bowel function, No. of patients	1492	585	415	2492		
Mean (95% CI)	94.0 (93.3-94.6)	93.4 (92.5-94.3)	94.0 (92.8-95.2)	93.8 (93.4-94.3)	.02	
Hormonal function, No. of patients	1467	563	412	2442		
Mean (95% CI)	89.8 (89.1-90.5)	86.7 (85.3-88.0)	89.7 (88.3-91.1)	89.1 (88.5-89.6)	<.001	

(continued)

Table.	Demograp	hics and	Baseline	Character	istics ((continued))
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	No. (%) of Patients				
	Radical Prostatectomy (n = 1523)	External Beam Radiation Therapy (n = 598)	Active Surveillance (n = 429)	Overall (N = 2550)	P Value
Medical Outcomes Study					
Short Form 36 score ^d					
Physical function, No. of patients	1477	577	405	2459	
Mean (95% CI)	87.9 (86.9-88.9)	78.3 (76.2-80.3)	84.0 (81.6-86.4)	85.0 (84.1-85.9)	<.001
Emotional well-being, No. of patients	1515	592	426	2533	
Mean (95% CI)	78.0 (77.1-78.9)	79.2 (77.7-80.7)	80.5 (78.9-82.1)	78.7 (78.0-79.4)	.10
Energy or fatigue, No. of patients	1477	577	405	2459	
Mean (95% CI)	72.4 (71.4-73.4)	68.3 (66.7-70.0)	71.5 (69.7-73.4)	71.3 (70.5-72.1)	<.001
Modified Social Support Scale score, No. of patients ^e	1515	592	426	2533	
Mean (95% CI)	81.2 (79.8-82.6)	79.0 (76.7-81.3)	80.5 (77.9-83.1)	80.6 (79.5-81.7)	.10
Center for Epidemiologic Studies score ^f					
Depression scale, No. of patients	1490	582	415	2487	
Mean (95% CI),	20.2 (19.3-21.2)	20.9 (19.2-22.5)	18.1 (16.4-20.0)	20.0 (19.3-20.8)	.12
Medical decision-making style, No. of patients ⁹	1504	585	411	2500	
Mean (95% CI)	78.7(77.7-79.7)	72.7 (70.7-74.7)	76.7 (74.4-79.0)	77.0 (76.1-77.8)	<.001
Accrual site, No. of patients	1523	598	429	2550	
Louisiana	392 (26)	219 (37)	108 (25)	719 (28)	
Utah	127 (8)	18 (3)	60 (14)	205 (8)	
Atlanta	195 (13)	55 (9)	58 (14)	308 (12)	
Los Angeles County, California	444 (29)	145 (24)	138 (32)	727 (29)	<.001
New Jersey	243 (16)	135 (23)	32 (7)	410 (16)	
Cancer of the Prostate Strategic Urologic Research Endeavor	122 (8)	26 (4)	33 (8)	181 (7)	

^a Based on the Total Illness Burden Index, scores range from 0 to 23, with higher score indicating greater severity and number of comorbid illnesses.

^b D'Amico risk classification system predicts risk of recurrence after treatment. Low-risk disease is defined as a clinical stage T2a or less, Gleason score 6 (3 + 3) or less, and a prostate-specific antigen (PSA) less than 10 ng/mL; high-risk disease, T2c or higher, Gleason score 8 (3 + 5, 4 + 4, 5 + 3) or greater, or a PSA value greater than 20 ng/mL; intermediate risk not defined by low- or high-risk definition.

^c Scores range from O to 100, with higher scores representing better function.

^d The Medical Outcomes Short-Form Health Survey 36 (SF-36) has 8 domains. The physical function domain score is a weighted sum of 10 items; emotional well-being, 5 items; energy or fatigue score, 4 items. Each domain score is directly transformed to a scale of 0 to 100 with increasing scores indicating better function or less disability.

^e Five questions were selected to create a modified domain score. Responses were directly transformed to a scale of 0 to 100 with increasing scores indicating greater support.

^f Derived using the 10-item Center for Epidemiologic Studies Depression Scale. Scores were scaled to 100, with higher scores indicating more severe depressive symptoms.

^g Seven items were scored on a scale from O to 100, with higher scores indicating increased patient choice, control, and responsibility.

Distinguishing between watchful waiting, active surveillance, and treatment delay was not possible. We categorized these patients as being actively surveilled recognizing that it was a heterogeneous group. For analysis, time 0 was the date patients underwent either radical prostatectomy or EBRT; whereas, the date of diagnosis was time 0 for patients who were being actively surveilled.

Statistical Analysis

Baseline characteristics were compared across treatments using Kruskal-Wallis tests for continuous variables and χ^2 tests for categorical variables.

To describe typical trajectories of function over time, longitudinal regression models were fit to predict EPIC domain scores as a function of treatment, time since treatment, and their interaction. For each domain, a single model was fit incorporating domain scores from all time points. We used generalized estimating equations (GEE) with an independent weight matrix because of the correlation between observations on the same patients. Modeling time using regression splines allowed for a flexible relationship between function and time. Variability in the interval between treatment and survey completion allowed for estimation of domain scores between rounds of data collection, and beyond 36 months.

Recognizing that outcomes (and patients' priorities) may differ by baseline function, we repeated these models, stratifying by baseline domain scores (excellent and less than excellent). Because excellent function has not been defined in the literature based on EPIC domain scores, a cutoff baseline score was selected for each domain that approximated the highest quartile of domain scores, an approach that has been used in prior publications on patient-reported outcomes after prostate cancer treatment.²⁰

To measure the association between treatment choice and domain score over time, a similar set of models was fit that adjusted for age, race/ethnicity, comorbidity,²¹ prostate cancer risk stratum,²² physical function,^{17,18,23} social support,²⁴ depression,²⁵ medical decision-making style,²⁶ site, and baseline EPIC domain score. This multivariable modeling approach was designed to minimize bias associated with known differences in baseline characteristics that are associated with functional outcomes (ie, confounding). Multiple imputation was used for missing covariates (see eMethods in the Supplement). Because androgen deprivation therapy is a standard component of radiation therapy for high-risk disease and an option in intermediate-risk disease, androgen deprivation therapy was not controlled for in the models.²⁷ Instead, exploratory models were fit for sexual and hormonal function with 5 treatment groups: nerve-sparing radical prostatectomy, non-nerve-sparing radical prostatectomy, EBRT without androgen deprivation therapy, EBRT with androgen deprivation therapy, and active surveillance. Unadjusted and adjusted longitudinal regression models using GEE were fit for responses to individual EPIC items and for the 3 SF-36 domains, using the same covariates as above. In the SF-36 regression models, the baseline SF-36 domain score was added as an independent variable.

Probability of overall and disease-specific survival was estimated by treatment using the Kaplan-Meier technique with log-rank tests.

Differences in domain scores between treatments were statistically significant if the 2-tailed *P* value was <.05. Domain scores were interpreted as clinically meaningful if the differences were as large as the MCID. R version 3.2.2 was used for all analyses.

Results

The parent study accrued 3709 patients, of whom 440 were excluded for failing to meet basic inclusion criteria. An additional 519 were excluded from the current study for receiving a treatment other than radical prostatectomy, EBRT, or active surveillance, leaving 2750 patients for consideration (eFigure 1 in the Supplement). The analytic cohort contained 2550 men (93%) who completed a baseline survey and at least 1 survey thereafter. Approximately 93% of surveys were completed on paper, while 7% were completed by telephone; 98% of surveys were conducted in English and 2% in Spanish; 54% of baseline surveys were collected prior to initial treatment. Survey response rates were 89% at 6 months, 86% at 12 months, and 78% at 36 months (eFigure 1 and eTable 1 in the Supplement).

Among men in the analytic cohort, 1523 (59.7%) underwent radical prostatectomy, 598 (23.5%) EBRT, and 429 (16.8%) active surveillance. Baseline characteristics are shown in the Table. Briefly, 26% of the cohort was nonwhite. Patients treated with EBRT were older, had higher comorbidity burden, and had higher-risk disease features than did patients who were treated with radical prostatectomy. Seventy-seven percent of active surveillance patients had low-risk disease.

Of the 1302 men (71%) who underwent radical prostatectomy and had complete reporting of nerve-sparing status, 859 (79%) had bilateral nerve-sparing, and of the 1032 (85%) who had complete reporting of the surgical approach, 1002 (77%) had bilateral nerve-sparing surgery. Of the 593 patients (99%) treated with EBRT who had complete reporting of utilization of androgen deprivation therapy, 265 (45%) received androgen deprivation therapy within the first year diagnosis of treatment; 378 patients (81%) of the 467 with complete records underwent IMRT. By the 3-year survey, 24.2% of active surveillance patients had undergone treatment, and 90.2% of the remainder had their PSA checked within the past 12 months.

For the stratified analyses, 26.4% of patients had excellent baseline domain scores for sexual function (≥90 points), 26.1% for urinary irritative symptoms (100 points), 61.7% for bowel function (100 points), and 39.1% for hormonal function (100 points).

Sexual Function

Patients undergoing radical prostatectomy had higher baseline sexual domain scores than men undergoing EBRT and had scores comparable with those on active surveillance (eTable 2 in the Supplement). Radical prostatectomy and EBRT were associated with declines in sexual function scores, but the decline was greater for patients who underwent radical prostatectomy, resulting in similar average unadjusted domain score for radical prostatectomy and EBRT at 3 years (Figure 1A, C, and D). The difference in functional decline between radical prostatectomy and EBRT was greater for the 26.4% of men with excellent baseline function, while the 73.6% of men with lower baseline function had poor sexual function outcomes regardless of whether they underwent radical prostatectomy or EBRT. Active surveillance was associated with preservation of function, with mild decline over time.

When controlling for baseline domain scores and other covariates (eTable 2; Figure 1B), men who underwent radical prostatectomy had a larger decline in sexual domain score than did those who underwent EBRT (adjusted mean domain score difference at 3 years, -11.9 points; 95% CI, -15.1 to -8.7) or active surveillance (-16.2; 95% CI, -20.6 to -11.7), relative to the MCID of 10 to 12 points. The adjusted mean domain score after EBRT was significantly worse than it was for active surveillance at 12 months (-10.5; 95% CI, -14.0 to -6.9), but the magnitude of difference at 3 years was no longer significant (-4.3; 95% CI, -9.2 to 0.7). Treatment, baseline domain score, and time since treatment were the only variables for which the magnitude of association with 3-year domain scores exceeded the MCID.

On exploratory analysis with a 5-tier treatment variable (nerve-sparing radical prostatectomy, non-nerve-sparing radical prostatectomy, EBRT alone, EBRT plus androgen deprivation therapy, and active surveillance), the mean difference between EBRT alone and active surveillance was not statistically significant (-3.0 points, P = .27), and the mean difference





C Men with excellent baseline domain score (≥90 points)

12

88

77

18

Time Since Treatment Start, mo

24

30



No. of patients

Active surveillance External beam radiation 558 Radical prostatectomy 1447 1407 1401

100

Sexual Function Domain Score, Unadjusted Mean (95% Cl) ~ 2 8







Active surveillance 97 External beam radiation 82



0.

Baseline

6

91

77



F Sexual function bother (reporting no, very small, or small problem)



See the Methods section and the Table footnotes for definitions measures, scoring ranges, and minimum clinically important differences. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. The number of patients represent those who completed surveys after enrollment. Shaded areas indicate 95% CIs.

556

1413

540

1395

B, Data in the forest plots are estimated by multivariable regression models that controlled for baseline domain score, age, race, comorbidity, prostate cancer risk group, physical function, social support, depression, medical decision-making style, and accrual site. Positive values represent better outcomes for the nonreferent group. eTable 8 in the Supplement contains unadjusted domain scores and number of patients for each subgroup by treatment. See the Table footnotes for definition of disease risk.

No. of patients

External beam radiation 567

Radical prostatectomy 1462

460

1284

between radical prostatectomy and EBRT plus androgen deprivation therapy was attenuated (-8.2 points; 95% CI, -13.2 to -3.2) lower than the MCID (eFigure 2 in the Supplement).

More men who underwent radical prostatectomy were bothered by sexual dysfunction 3 years after diagnosis (44% vs 35% for EBRT and 28% for active surveillance, P < .001 on multivariable analysis; Figure 1F; eTable 2 in the Supplement). Erection insufficient for intercourse was common at 3 years (70% for radical prostatectomy, 71% for EBRT, and 51% for active surveillance on raw percentages, Figure 1E), but when controlling for baseline sexual function and other factors, the odds were significantly higher for radical prostatectomy than for active surveillance (odds ratio [OR], 3.4; 95% CI, 2.5 to 4.6) and for radical prostatectomy than EBRT (OR, 2.1; 95% CI, 1.5 to 2.9). Among men who had sufficient erections at baseline, erection sufficient for intercourse at 3 years was reported in 43% (95% CI, 40% to 47%) of men who had undergone radical prostatectomy; 53% (95% CI, 45% to 60%), EBRT; and 75% (95% CI, 68%-80%), active surveillance, in raw percentages. An exploratory multivariable model, using 5 treatment groups, yielded similar results (eTable 3 in the Supplement).

Urinary Incontinence

Baseline urinary incontinence domain scores were similar across groups (eTable 4 in the Supplement). However, radical prostatectomy was associated with a significant decline in urinary incontinence score after treatment, particularly in the 60.3% of men with perfect urinary incontinence domain scores at baseline (Figure 2A, C, and D). There was no significant change in urinary incontinence score for men who had EBRT or active surveillance, regardless of baseline score.

Despite some improvement in incontinence domain scores 12 months after radical prostatectomy, adjusted mean incontinence scores were still significantly worse for radical prostatectomy than for active surveillance (–12.7 points, 95% CI, –16.0 to –9.3) and EBRT (–18.0 points, 95% CI, –20.5 to –15.4) at 3 years, differences greater than the MCID (6 points) (Figure 2B; eTable 4 in the Supplement). By contrast, urinary incontinence was not significantly different between EBRT and active surveillance. Treatment, baseline domain score, and time since treatment were the only variables for which the magnitude of association with the 3-year domain score exceeded the MCID.

Reports of moderate or big problems with urinary leakage were more common after radical prostatectomy vs active surveillance (14% vs 6%; OR, 2.9; 95% CI, 1.8-4.7) and radical prostatectomy than EBRT (14% vs 5%; OR, 4.5; 95% CI, 2.7-7.3; Figure 2E; and eTable 4 in the Supplement). Urinary function bother scores were not significantly different for radical prostatectomy vs active surveillance and EBRT vs active surveillance at 3 years but were higher for radical prostatectomy vs EBRT (12% vs 10%; OR, 1.7; 1.1-2.5; Figure 2F; and eTable 4 in the Supplement).

Urinary Irritative Symptoms

Baseline scores were similar across groups (eTable 4 in the Supplement). Scores improved for radical prostatectomy, particularly for the 73.9% of men whose baseline score was

less than 100 (**Figure 3**A, C, and D). Those undergoing EBRT or active surveillance experienced little or no change in irritative urinary symptoms.

Adjusted urinary irritative function scores were slightly better for men undergoing radical prostatectomy than being actively surveilled at 1 year (4.5 points; 95% CI, 3.0-6.0) and 3 years (5.2 points, 95% CI, 3.2-7.2]), at the threshold of clinical significance (Figure 3B; eTable 4 in the Supplement). Other comparisons across treatments, while statistically significant, were lower than the MCID of 5 (Figure 3B; eTable 4 in the Supplement). Besides treatment with radical prostatectomy, the only other factors for which the magnitude of association with 3-year domain score exceeded the MCID were baseline domain score and time since treatment.

Reports of moderate or big problems with burning with urination were uncommon (2% in each group; Figure 3E; eTable 4 in the Supplement). Reports of moderate or big problem with frequent urination were lower for radical prostatectomy than for active surveillance (13% vs 18%; OR, 0.6; 95% CI, 0.4-0.8]) and for EBRT vs active surveillance (15% vs 18%, OR, 0.6; 95%, 0.4-0.8) at 3 years, but not significantly different between radical prostatectomy and EBRT (Figure 3F; eTable 4 in the Supplement).

Bowel Function

Decline in bowel domain score was not common (Figure 4A, C, and D; eTable 5 in the Supplement). Six months after treatment, the mean domain scores were higher in men who underwent radical prostatectomy than who underwent EBRT (4.6 points, 95% CI, 3.2-6.1) and lower for EBRT vs active surveillance (-5.8 points; 95% CI, -10.3 to -1.2 points). However, by 12 months these differences were near the MCID of 4 and by 36 months, they were smaller. Unadjusted and adjusted results were similar (Figure 4B). No other independent variables had a magnitude of association with 3-year domain score that met the threshold for clinical significance.

The frequency of moderate or big problem with bowel bother, bloody stools, or bowel urgency was 1% to 8% across all treatments at 3 years (Figure 4E-F; eTable 5 in the Supplement). Nevertheless, the odds of bowel urgency at 3 years were lower for radical prostatectomy than EBRT (3% vs 7%, OR, O.3; 95% CI, 0.2-0.6) and radical prostatectomy vs active surveillance (3% vs 5%, OR, 0.5; 95%, 0.3-0.9).

Hormone Function

The mean hormone domain scores were worse for EBRT than for active surveillance and radical prostatectomy at 6 months (radical prostatectomy vs EBRT, 5.0 points; 95% CI, 3.3 to 6.6 points; EBRT vs active surveillance, -6.5 points; 95% CI, -11.1 to -1.9), but these differences no longer significant at 3 years on unadjusted or adjusted analyses (**Figure 5**; eTable 6 in the **Supplement**). No other independent variables had a magnitude of association with 3-year domain score that reached the MCID.

In the exploratory models that separated EBRT into with and without androgen deprivation therapy, the only group with decrements in hormone function was the EBRT plus





E	Baseline	6	12	18	24	30	36	Baseline	6	12	18	24
			Time Since	Treatmen	t Start, mo					Time Since	Treatmen	t Start, mo
No. of patients												
Active surveillance	415	410	385				347	405	409	376		
External beam radiation	582	579	551				474	580	576	543		
Radical prostatectomy	1492	1442	1411				1293	1468	1432	1394		

See the Methods section for explanation of the measures, scoring ranges, and minimum clinically important difference definitions. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. The

number of patients indicate those who completed surveys at the specified time following enrollment. Shaded areas indicate 95% Cls. See Figure 1 legend for an explanation of the forest plot.

No.

30

36

345 477

1302

Figure 3. Association Between Initial Treatment of Prostate Cancer and Urinary Irritative Outcomes









D Men with lower baseline domain score (<100 points)



See the Methods section for explanation of the measures, scoring ranges, and minimum clinically important difference definitions. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. The

409

576

1439

385

554

1410

number of patients indicate those who completed surveys at the specified time following enrollment. Shaded areas indicate 95% Cis. See Figure 1 legend for an explanation of the forest plot.

382

555

1413

409

577

1440

414

585

1491

413

No. of patients Active surveillance

External beam radiation 582

Radical prostatectomy 1488

347

479

1299

347

477

1298





No. of patients Active surveillance External beam radiation 331 Radical prostatectomy



F Bowel function bother (reporting no, very small, or small problem)



See the Methods section for explanation of the measures, scoring ranges, and minimum clinically important difference definitions. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. The

number of patients indicate those who completed surveys at the specified time following enrollment. Shaded areas indicate 95% Cls. See Figure 1 legend for an explanation of the forest plot.











See the Methods section for explanation of the measures, scoring ranges, and minimum clinically important difference definitions. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. The

number of patients indicate those who completed surveys at the specified time following enrollment. Shaded areas indicate 95% CIs. See Figure 1 legend for an explanation of the forest plot.

were 78 deaths, including 3 prostate cancer deaths. On

Kaplan-Meier analysis, estimated 3-year disease-specific survival was not significantly different across groups (99.9% for radical prostectomy, 99.7% for EBRT, and 100%

for active surveillance). Unadjusted 3-year overall survival

was higher for radical prostatectomy (99%, 95% CI,

98%-99%]) than for other groups (EBRT, 96%; 95% CI, 94%-

98%; active surveillance, 97%; 95% CI, 95-99; P < .001),

androgen deprivation therapy group, and these associations were limited to the first year (eFigure 2 in the Supplement).

Health-related Quality of Life

Baseline physical functioning and energy or fatigue scores on the SF-36 were lower for men undergoing EBRT than radical prostatectomy or active surveillance (eTable 7 in the Supplement). None of the treatment groups experienced a clinically significant decline in physical functioning, emotional well-being, or energy or fatigue scores (Figure 6). On multivariable analysis, associations between treatment and 3-year SF-36 qualityof-life domain scores were below the threshold for clinical significance, as were associations baseline EPIC sexual and urinary incontinence domain scores and 3-year SF-36 domain scores.

Survival

Median follow-up time among censored patients was 40 months (interquartile range [IQR], 38-45, months). There

commensurate with the younger age and lower comorbidity of men undergoing radical prostatectomy (eTable 9 in the Supplement). Discussion

> In this study of men with localized prostate cancer, radical prostatectomy was associated with clinically significant declines in sexual function compared with EBRT and active





See the Methods section for explanation of the Medical Outcomes Study measures, scoring ranges, and minimum clinically important difference definitions. Baseline is defined as the date of initiation of treatment (in men undergoing radical prostatectomy or radiation) or date of diagnosis for those who were actively surveilled. Solid lines indicate unadjusted means; shaded areas, 95% CIs. Shaded areas indicate 95% CIs. See Figure 1 legend for an explanation of the forest plot.

surveillance, particularly among men with excellent function at baseline. Urinary incontinence scores also declined significantly after surgery compared with EBRT and active surveillance, with 14% of patients treated with radical prostatectomy reporting a moderate or big problem with urinary leakage at 3 years compared with 5% with EBRT and 6% with active surveillance. Radical prostatectomy was associated with better irritative voiding symptoms than active surveillance, with a difference that met the threshold for clinical significance. Mean scores in bowel and hormonal domains were significantly worse for EBRT vs radical prostatectomy and active surveillance at 6 months, but the differences were below threshold for clinical significance by 3 years. Treatment, baseline domain scores, and time since treatment were the independent variables with clinically significant associations with 3-year domain scores. None of the treatment groups experienced clinically significant declines in health-related quality-of-life domain scores. This information may facilitate patient counseling regarding the expected harms of contemporary treatments and their possible effect on quality of life.

Prior studies have quantified the harms of prostate cancer treatment. However, randomized trials studying localized prostate cancer have been difficult to execute, and those that have been completed focused on outmoded treatments; enrolled too few minority patients; lacked a range of disease severity; failed to collect baseline functional assessments; or included a preponderance of elderly, infirm, and lowrisk patients, for whom treatment is questionable.^{3,5,6,28-30} The ProtecT trial,^{5,6} for example, included 99% white patients and nearly 80% of patients with a Gleason score of 6 (low-risk). In ProtecT, 87% of surgical patients underwent open radical prostatectomy (vs 77% who underwent robotic surgery in this study) and patients undergoing EBRT had 3-dimentional conformal radiation therapy plus androgen deprivation therapy (compared with 81% receiving IMRT, with 45% receiving concurrent androgen deprivation therapy in this study). Thus, the ProtecT study findings may be difficult to apply to a racially diverse population with a range of disease risk strata, managed with contemporary treatments.

Case series that have evaluated functional outcomes are not generalizable because they reported on outcomes at centers of excellence; lacked the variables necessary to adjust for confounding; lacked an active surveillance group as a comparator; or had other sources of bias.³¹⁻³⁷

Despite these caveats, functional outcomes in this study are similar to previously published multi-institutional prospective cohort studies and to the ProtecT trial.^{6,20,38-41} Nevertheless, comparisons between the CEASAR cohort and similar historical cohorts have shown slightly smaller declines in erectile function domain scores at 6 and 12 months with robotic radical prostatectomy than with open radical prostatectomy, and slightly better bowel domain scores at 6 months for IMRT than for older 3-dimensional conformal radiation therapy.^{42,43} These data suggest that contemporary treatments have similar associations with functional outcomes but perhaps slightly less in magnitude. This study may have implications for decision making for patients with localized prostate cancer. First, it demonstrates the frequency and severity of adverse effects of contemporary treatments and the likelihood of preserved global quality of life regardless of treatment, thus providing a basis for shared decision making. Second, in contrast to previously published studies, this study may be more generalizable, since the cohort is racially diverse, population based, and includes a range of disease severity.^{3,6,28,38} Third, this study may inform future research on personalized risk assessment, tools to facilitate shared decision making, and other patient-centered outcomes.

Limitations

This study has several limitations. There may be disagreement about the definition of MCID, which may also differ from one patient to the next. Although some outcomes favored one treatment over another, the results do not indicate what value patients place on particular domains. Furthermore, there are other important outcomes to consider in localized prostate cancer, including long-term functional outcomes and oncologic end points, anxiety, satisfaction, and financial toxicity. The number and severity of adverse outcomes presenting beyond 3 years may differ by treatment, and 3 years is inadequate to estimate oncologic outcomes. Data on patients who had other treatments, such as brachytherapy and ablation, were not included because there were not enough patients who received these treatments to generate sufficient statistical power for reliable comparisons. Aggregated data and average function scores may fail to capture the severity of adverse effects for individuals and do not yield personalized risk estimates. The analysis did not adjust for the quality of care or experience of the treating clinician or institution, which may influence outcomes. Thus, the findings of this study represent a subset of the information needed to guide decision making. A substantial proportion of patients answered the baseline survey after initiating treatment, raising the possibility of recall bias, although in prior studies the magnitude of recall bias was small for the EPIC survey.¹⁶ This study used an observational cohort, rather than an experimental design, so there may be unmeasured sources of confounding.

Conclusions

In this cohort of men with localized prostate cancer, radical prostatectomy was associated with a greater decrease in sexual function and urinary incontinence than either EBRT or active surveillance after 3 years and was associated with fewer urinary irritative symptoms than active surveillance; however, no meaningful differences existed in either bowel or hormonal function beyond 12 months or in other domains of health-related quality of life measures. These findings may facilitate counseling regarding the comparative harms of contemporary treatments for prostate cancer.

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REFERENCES

1. Thompson I, Thrasher JB, Aus G, et al; AUA Prostate Cancer Clinical Guideline Update Panel. Guideline for the management of clinically localized prostate cancer: 2007 update. *J Urol*. 2007;177(6): 2106-2131.

2. Makarov D, Fagerlin A, Chrouser K, et al. AUA white paper on implementation of shared decision making into urological practice. *Urol Pract.* 2015;3 (5):1-31.

3. Wilt TJ, Brawer MK, Jones KM, et al; Prostate Cancer Intervention versus Observation Trial (PIVOT) Study Group. Radical prostatectomy versus observation for localized prostate cancer. *N Engl J Med*. 2012;367(3):203-213.

4. Bill-Axelson A, Holmberg L, Garmo H, et al. Radical prostatectomy or watchful waiting in early prostate cancer. *N Engl J Med*. 2014;370(10):932-942.

5. Hamdy FC, Donovan JL, Lane JA, et al 10-Year outcomes after monitoring, surgery, or radiotherapy for localized prostate cancer. *N Engl J Med*. 2016:375(15):1415-1424.

6. Donovan JL, Hamdy FC, Lane JA, et al; ProtecT Study Group. Patient-reported outcomes after monitoring, surgery, or radiotherapy for prostate cancer. *N Engl J Med*. 2016;375(15):1425-1437.

7. Sheets NC, Goldin GH, Meyer AM, et al. Intensity-modulated radiation therapy, proton therapy, or conformal radiation therapy and morbidity and disease control in localized prostate cancer. *JAMA*. 2012;307(15):1611-1620.

8. Jacobs BL, Zhang Y, Skolarus TA, et al. Managed care and the diffusion of intensity-modulated radiotherapy for prostate cancer. *Urology*. 2012;80 (6):1236-1242.

9. Jacobs BL, Zhang Y, Skolarus TA, Hollenbeck BK. Growth of high-cost intensity-modulated radiotherapy for prostate cancer raises concerns about overuse. *Health Aff (Millwood)*. 2012;31(4): 750-759.

10. Yanamadala S, Chung BI, Hernandez-Boussard TM. Robot-assisted versus open radical

prostatectomy utilization in hospitals offering robotics. *Can J Urol*. 2016;23(3):8279-8284.

11. Oberlin DT, Flum AS, Lai JD, Meeks JJ. The effect of minimally invasive prostatectomy on practice patterns of American urologists. *Urol Oncol*. 2016; 34(6):255.e1-255.e5.

12. Lowrance WT, Eastham JA, Savage C, et al. Contemporary open and robotic radical prostatectomy practice patterns among urologists in the United States. *J Urol*. 2012;187(6):2087-2092.

13. Barocas DA, Chen V, Cooperberg M, et al. Using a population-based observational cohort study to address difficult comparative effectiveness research questions: the CEASAR study. *J Comp Eff Res.* 2013;2(4):445-460.

14. Szymanski KM, Wei JT, Dunn RL, Sanda MG. Development and validation of an abbreviated version of the expanded prostate cancer index composite instrument for measuring health-related quality of life among prostate cancer survivors. *Urology*. 2010;76(5):1245-1250.

 Skolarus TA, Dunn RL, Sanda MG, et al; PROSTQA Consortium. Minimally important difference for the expanded prostate cancer index composite short form. *Urology*. 2015;85(1):101-105.

16. Resnick MJ, Barocas DA, Morgans AK, et al. Contemporary prevalence of pretreatment urinary, sexual, hormonal, and bowel dysfunction: defining the population at risk for harms of prostate cancer treatment. *Cancer*. 2014;120(8):1263-1271.

17. Ware JE Jr, Sherbourne CD. The MOS 36-item short-form health survey (SF-36), I: conceptual framework and item selection. *Med Care*. 1992;30 (6):473-483.

18. McHorney CA, Ware JE Jr, Raczek AE. The MOS 36-Item Short-Form Health Survey (SF-36), II: psychometric and clinical tests of validity in measuring physical and mental health constructs. *Med Care*. 1993;31(3):247-263.

19. Jayadevappa R, Malkowicz SB, Wittink M, Wein AJ, Chhatre S. Comparison of distribution- and anchor-based approaches to infer changes in health-related quality of life of prostate cancer survivors. *Health Serv Res.* 2012;47(5):1902-1925.

20. Resnick MJ, Koyama T, Fan K-H, et al. Long-term functional outcomes after treatment for localized prostate cancer. *N Engl J Med.* 2013;368 (5):436-445.

21. Stier DM, Greenfield S, Lubeck DP, et al. Quantifying comorbidity in a disease-specific cohort: adaptation of the total illness burden index to prostate cancer. *Urology*. 1999;54(3):424-429.

22. D'Amico AV, Whittington R, Malkowicz SB, et al. Pretreatment nomogram for prostate-specific antigen recurrence after radical prostatectomy or external-beam radiation therapy for clinically localized prostate cancer. *J Clin Oncol*. 1999;17(1): 168-172.

23. Haley SM, McHorney CA, Ware JE Jr. Evaluation of the MOS SF-36 physical functioning scale (PF-10), I: Unidimensionality and reproducibility of the Rasch item scale. *J Clin Epidemiol*. 1994;47(6): 671-684.

24. Sherbourne CD, Stewart AL. The MOS social support survey. *Soc Sci Med*. 1991;32(6):705-714.

25. Andresen EM, Malmgren JA, Carter WB, Patrick DL. Screening for depression in well older adults:

Research Original Investigation

evaluation of a short form of the CES-D (Center for Epidemiologic Studies Depression Scale). *Am J Prev Med.* 1994;10(2):77-84.

26. Kaplan SH, Greenfield S, Gandek B, Rogers WH, Ware JE Jr. Characteristics of physicians with participatory decision-making styles. *Ann Intern Med*. 1996;124(5):497-504.

27. Mohler J, Bahnson RR, Boston B, et al. NCCN clinical practice guidelines in oncology: prostate cancer. J Natl Compr Canc Netw. 2010;8(2):162-200.

28. Johansson E, Steineck G, Holmberg L, et al; SPCG-4 Investigators. Long-term quality-of-life outcomes after radical prostatectomy or watchful waiting: the Scandinavian Prostate Cancer Group-4 randomised trial. *Lancet Oncol.* 2011;12(9):891-899.

29. Bill-Axelson A, Holmberg L, Garmo H, et al. Radical prostatectomy or watchful waiting in early prostate cancer. *N Engl J Med*. 2014;370(10):932-942.

30. Yaxley JW, Coughlin GD, Chambers SK, et al. Robot-assisted laparoscopic prostatectomy versus open radical retropubic prostatectomy: early outcomes from a randomised controlled phase 3 study. *Lancet*. 2016;388(10049):1057-1066.

31. Woo SH, Kang DI, Ha YS, et al. Comprehensive analysis of sexual function outcome in prostate cancer patients after robot-assisted radical prostatectomy. *J Endourol.* 2014;28(2):172-177.

32. Hollenbeck BK, Dunn RL, Wei JT, Montie JE, Sanda MG. Determinants of long-term sexual health outcome after radical prostatectomy measured by a validated instrument. *J Urol.* 2003;169(4):1453-1457.

33. Kundu SD, Roehl KA, Eggener SE, Antenor JA, Han M, Catalona WJ. Potency, continence and complications in 3,477 consecutive radical retropubic prostatectomies. *J Urol*. 2004;172(6 pt 1):2227-2231.

34. Zelefsky MJ, Poon BY, Eastham J, Vickers A, Pei X, Scardino PT. Longitudinal assessment of quality of life after surgery, conformal brachytherapy, and intensity-modulated radiation therapy for prostate cancer. *Radiother Oncol.* 2016;118(1):85-91.

35. Ficarra V, Novara G, Ahlering TE, et al. Systematic review and meta-analysis of studies reporting potency rates after robot-assisted radical prostatectomy. *Eur Urol*. 2012;62(3):418-430. doi: 10.1016/j.eururo.2012.05.046

36. Ficarra V, Novara G, Rosen RC, et al. Systematic review and meta-analysis of studies reporting urinary continence recovery after robot-assisted radical prostatectomy. *Eur Urol.* 2012;62(3):405-417. doi:10.1016/j.eururo.2012.05.045

37. Bourke L, Boorjian SA, Briganti A, et al. Survivorship and improving quality of life in men with prostate cancer. *Eur Urol*. 2015;68(3):374-383. **38**. Sanda MG, Dunn RL, Michalski J, et al. Quality of life and satisfaction with outcome among prostate-cancer survivors. *N Engl J Med*. 2008;358 (12):1250-1261.

39. Alemozaffar M, Regan MM, Cooperberg MR, et al. Prediction of erectile function following treatment for prostate cancer. *JAMA*. 2011;306(11): 1205-1214.

40. Stanford JL, Feng Z, Hamilton AS, et al. Urinary and sexual function after radical prostatectomy for clinically localized prostate cancer: the Prostate Cancer Outcomes Study. *JAMA*. 2000;283(3):354-360.

41. Punnen S, Cowan JE, Chan JM, Carroll PR, Cooperberg MR. Long-term health-related quality of life after primary treatment for localized prostate cancer: results from the CaPSURE registry. *Eur Urol*. 2015;68(4):600-608.

42. O'Neil B, Koyama T, Alvarez J, et al. The comparative harms of open and robotic prostatectomy in population based samples. *J Urol*. 2016;195(2):321-329.

43. O'Neil B, Hoffman KE, Koyama T, et al. Population-based comparison of patient-reported function after 3-dimensional conformal vs contemporary external beam radiation therapy. *Int J Radiat Oncol Biol Phys*. 2016;96(25):E399.